

Supplement to the paper “An asymptotic equivalence of the cross-data and predictive estimators”

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This article supplements Ogasawara (2019).

S1. Bernoulli distribution

S1.1 Bernoulli distribution under canonical parametrization

$$\begin{aligned}
 \Pr(x | \theta) &= \left\{ \frac{1}{1 + \exp(-\theta)} \right\}^x \left\{ \frac{\exp(-\theta)}{1 + \exp(-\theta)} \right\}^{1-x} \quad (x = 0, 1), \\
 &= \frac{\exp\{-(1-x)\theta\}}{1 + \exp(-\theta)}, \quad \left(\hat{\theta}_{\text{ML}(-i)} = \ln \frac{\bar{x}_{(-i)}}{1 - \bar{x}_{(-i)}} \right), \\
 D_{(-)} &= \sum_{i=1}^n \left[(1 - x_i) \omega \hat{\theta}_{\text{ML}(-i)} + \ln \{1 + \exp(-\omega \hat{\theta}_{\text{ML}(-i)})\} \right], \\
 \frac{\partial D_{(-)}}{\partial \omega} &= \sum_{i=1}^n \left\{ (1 - x_i) \hat{\theta}_{\text{ML}(-i)} - \frac{\exp(-\omega \hat{\theta}_{\text{ML}(-i)})}{1 + \exp(-\omega \hat{\theta}_{\text{ML}(-i)})} \hat{\theta}_{\text{ML}(-i)} \right\}, \\
 &= \sum_{i=1}^n \left\{ 1 - x_i - \frac{\{\bar{x}_{(-i)} / (1 - \bar{x}_{(-i)})\}^{-\omega}}{1 + \{\bar{x}_{(-i)} / (1 - \bar{x}_{(-i)})\}^{-\omega}} \right\} \ln \frac{\bar{x}_{(-i)}}{1 - \bar{x}_{(-i)}} \\
 &= \sum_{i=1}^n \left\{ \frac{1}{1 + \{\bar{x}_{(-i)} / (1 - \bar{x}_{(-i)})\}^{-\omega}} - x_i \right\} \ln \frac{\bar{x}_{(-i)}}{1 - \bar{x}_{(-i)}}, \tag{S1.1}
 \end{aligned}$$

where

$$\bar{x}_{(-i)} = \bar{x} - (n-1)^{-1} (x_i - \bar{x}) \quad (i = 1, \dots, n). \tag{S1.2}$$

To solve $\partial D_{(-)} / \partial \omega = 0$, we use

$$\begin{aligned}
\frac{\partial^2 D_{(-)}}{\partial \omega^2} &= \sum_{i=1}^n \frac{\exp(-\omega \hat{\theta}_{\text{ML}(-i)})}{\{1 + \exp(-\omega \hat{\theta}_{\text{ML}(-i)})\}^2} \hat{\theta}_{\text{ML}(-i)}^2 \\
&= \sum_{i=1}^n \frac{\{\bar{x}_{(-i)} / (1 - \bar{x}_{(-i)})\}^{-\omega}}{[1 + \{\bar{x}_{(-i)} / (1 - \bar{x}_{(-i)})\}^{-\omega}]^2} \left(\ln \frac{\bar{x}_{(-i)}}{1 - \bar{x}_{(-i)}} \right)^2. \tag{S1.3}
\end{aligned}$$

S1.2 Bernoulli distribution using the expectation parameter

$$\Pr(x | \pi) = \pi^x (1 - \pi)^{1-x} \quad (x = 0, 1), \quad (\hat{\pi}_{\text{ML}(-i)} = \bar{x}_{(-i)}),$$

$$D_{(-)} = \sum_{i=1}^n \{-x_i \ln(\omega \hat{\pi}_{\text{ML}(-i)}) - (1 - x_i) \ln(1 - \omega \hat{\pi}_{\text{ML}(-i)})\}$$

$$= \sum_{i=1}^n \{-x_i \ln(\omega \bar{x}_{(-i)}) - (1 - x_i) \ln(1 - \omega \bar{x}_{(-i)})\},$$

$$\frac{\partial D_{(-)}}{\partial \omega} = \sum_{i=1}^n \left\{ -\frac{x_i}{\omega} + \frac{(1 - x_i)\bar{x}_{(-i)}}{1 - \omega \bar{x}_{(-i)}} \right\} = -\frac{n \bar{x}}{\omega} + \sum_{i=1}^n \frac{(1 - x_i)\bar{x}_{(-i)}}{1 - \omega \bar{x}_{(-i)}}, \tag{S1.4}$$

$$\frac{\partial^2 D_{(-)}}{\partial \omega^2} = \frac{n \bar{x}}{\omega^2} + \sum_{i=1}^n \frac{(1 - x_i)\bar{x}_{(-i)}^2}{(1 - \omega \bar{x}_{(-i)})^2}.$$

S2. Gamma distribution when the shape parameter α and the scale parameter β are unknown

$$f(x | \alpha, \beta) = \frac{x^{\alpha-1} \exp(-x/\beta)}{\beta^\alpha \Gamma(\alpha)} \quad (x > 0), \quad \boldsymbol{\omega} = (\omega_\beta, \omega_\alpha)',$$

$$\begin{aligned}
D_{(-)} &= \sum_{i=1}^n \left\{ -(\omega_\alpha \hat{\alpha}_{\text{ML}(-i)} - 1) \ln x_i + \omega_\alpha \hat{\alpha}_{\text{ML}(-i)} \ln(\omega_\beta \hat{\beta}_{\text{ML}(-i)}) \right. \\
&\quad \left. + \frac{x_i}{\omega_\beta \hat{\beta}_{\text{ML}(-i)}} + \ln \Gamma(\omega_\alpha \hat{\alpha}_{\text{ML}(-i)}) \right\}, \tag{S2.1}
\end{aligned}$$

$$\begin{aligned}
\frac{\partial D_{(-)}}{\partial \omega_\beta} &= \sum_{i=1}^n \left(\frac{\omega_\alpha \hat{\alpha}_{\text{ML}(-i)}}{\omega_\beta} - \frac{x_i}{\omega_\beta^2 \hat{\beta}_{\text{ML}(-i)}} \right), \quad (\hat{\beta}_{\text{ML}(-i)} = \bar{x}_{(-i)} / \hat{\alpha}_{\text{ML}(-i)}), \\
&= \sum_{i=1}^n \left(\frac{\omega_\alpha \hat{\alpha}_{\text{ML}(-i)}}{\omega_\beta} - \frac{\hat{\alpha}_{\text{ML}(-i)} x_i}{\omega_\beta^2 \bar{x}_{(-i)}} \right), \\
\frac{\partial D_{(-)}}{\partial \omega_\alpha} &= \sum_{i=1}^n \{-\hat{\alpha}_{\text{ML}(-i)} \ln x_i + \hat{\alpha}_{\text{ML}(-i)} \ln(\omega_\beta \hat{\beta}_{\text{ML}(-i)}) + \psi(\omega_\alpha \hat{\alpha}_{\text{ML}(-i)}) \hat{\alpha}_{\text{ML}(-i)}\}, \\
(\psi(\alpha)) &= \partial \ln \Gamma(\alpha) / \partial \alpha.
\end{aligned}$$

From $\partial D_{(-)} / \partial \omega_\beta = 0$, we obtain

$$\hat{\omega}_\beta = \frac{\sum_{i=1}^n \hat{\alpha}_{\text{ML}(-i)} x_i / \bar{x}_{(-i)}}{\hat{\omega}_\alpha \sum_{i=1}^n \hat{\alpha}_{\text{ML}(-i)}}. \quad (\text{S2.2})$$

From (S2.2) with $\hat{\beta}_{\text{ML}(-i)} = \bar{x}_{(-i)} / \hat{\alpha}_{\text{ML}(-i)}$, we have

$$\begin{aligned}
\frac{\partial D_{(-)}}{\partial \omega_\alpha} &= \sum_{i=1}^n \left\{ -\hat{\alpha}_{\text{ML}(-i)} \ln x_i + \hat{\alpha}_{\text{ML}(-i)} \ln \frac{(\bar{x}_{(-i)} / \hat{\alpha}_{\text{ML}(-i)}) \sum_{j=1}^n \hat{\alpha}_{\text{ML}(-j)} x_j / \bar{x}_{(-j)}}{\omega_\alpha \sum_{j=1}^n \hat{\alpha}_{\text{ML}(-j)}} \right. \\
&\quad \left. + \psi(\omega_\alpha \hat{\alpha}_{\text{ML}(-i)}) \hat{\alpha}_{\text{ML}(-i)} \right\}, \\
\frac{\partial^2 D_{(-)}}{\partial \omega_\alpha^2} &= \sum_{i=1}^n \left\{ -\frac{\hat{\alpha}_{\text{ML}(-i)}}{\omega_\alpha} + \psi'(\omega_\alpha \hat{\alpha}_{\text{ML}(-i)}^2) \hat{\alpha}_{\text{ML}(-i)}^2 \right\}, \\
(\psi'(\alpha)) &= \partial \psi(\alpha) / \partial \alpha.
\end{aligned} \quad (\text{S2.3})$$

S3. Multivariate multiple regression

S3.1 Multivariate multiple regression with normal errors when regression coefficients and error variances/covariances are unknown

It is assumed that

$$\mathbf{y}_i \sim N(\mathbf{Bx}_i, \boldsymbol{\Sigma}) \quad (i = 1, \dots, n), \quad (\text{S3.1})$$

where \mathbf{y}_i is the $q^* \times 1$ response vector; \mathbf{B} is the $q^* \times p^*$ matrix of regression coefficients; \mathbf{x}_i is the $p^* \times 1$ vector of fixed covariates; and $\boldsymbol{\Sigma}$ is the $q^* \times q^*$ covariance matrix of \mathbf{y}_i .

For the cross-data estimators of \mathbf{B} and $\boldsymbol{\Sigma}$, we have
 $\boldsymbol{\omega} = \{\text{vec}'(\boldsymbol{\Omega}_B), \text{vec}'(\boldsymbol{\Omega}_\Sigma)\}',$

$$L_{(-)} \equiv \prod_{i=1}^n \frac{1}{(2\pi)^{q^*/2} |\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)}|^{1/2}} \exp\{-(\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' \\ \times (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i) / 2\},$$

$$D_{(-)} = -\ln L_{(-)} = \frac{nq^*}{2} \ln(2\pi) + \frac{1}{2} \sum_{i=1}^n \ln |\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)}| \\ + \frac{1}{2} \sum_{i=1}^n (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} \\ \times (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i),$$

$$\frac{\partial D_{(-)}}{\partial \omega_{Bab}} = -\sum_{i=1}^n \{(\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' \mathbf{x}_i'\}_{ab} \\ \times (\hat{\mathbf{B}}_{ML(-i)})_{ab},$$

$$(a = 1, \dots, q^*; b = 1, \dots, p^*),$$

$$\hat{\mathbf{B}}_{ML(-i)} = (\mathbf{X}_{(-i)}' \mathbf{X}_{(-i)})^{-1} \mathbf{X}_{(-i)}' \mathbf{Y}_{(-i)}, \quad (\text{S3.2})$$

where $\mathbf{X}_{(-i)} = (\mathbf{x}_1, \dots, \mathbf{x}_n)'$ with \mathbf{x}_i being deleted and $\mathbf{Y}_{(-i)}$ is similarly defined,

$$\hat{\boldsymbol{\Sigma}}_{ML(-i)} = (n-1)^{-1} (\mathbf{Y}_{(-i)}' - \hat{\mathbf{B}}_{ML(-i)} \mathbf{X}_{(-i)}') (\mathbf{Y}_{(-i)} - \mathbf{X}_{(-i)} \hat{\mathbf{B}}_{ML(-i)}'), \\ \frac{\partial D_{(-)}}{\partial \omega_{\Sigma ef}} = \frac{1}{2} \sum_{i=1}^n (2 - \delta_{ef}) \{(\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1}\}_{ef} \hat{\sigma}_{ML(-i)ef} \\ - \frac{1}{2} \sum_{i=1}^n (2 - \delta_{ef}) \{(\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)\}_e \\ \times \{(\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)\}_f \hat{\sigma}_{ML(-i)ef} \quad (\text{S3.3})$$

$$(q^* \geq e \geq f \geq 1),$$

$$\frac{\partial^2 D_{(-)}}{\partial \omega_{Bab} \partial \omega_{Bcd}} = \sum_{i=1}^n \{(\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1}\}_{ac} (\mathbf{x}_i \mathbf{x}_i')_{db} (\hat{\mathbf{B}}_{ML(-i)})_{ab} (\hat{\mathbf{B}}_{ML(-i)})_{cd}$$

$$(a, c = 1, \dots, q^*; b, d = 1, \dots, p^*),$$

$$\begin{aligned}
\frac{\partial^2 D_{(-)}}{\partial \omega_{Bab} \partial \omega_{\Sigma ef}} &= \sum_{i=1}^n \sum_{(e,f)}^2 \frac{1}{2} (2 - \delta_{ef}) \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} \}_{ae} \\
&\times \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' \}_{fb} (\hat{\mathbf{B}}_{ML(-i)})_{ab} \hat{\sigma}_{ML(-i)ef} \\
(a = 1, \dots, q^*; b = 1, \dots, p^*; q^* \geq e \geq f \geq 1), \\
\frac{\partial^2 D_{(-)}}{\partial \omega_{\Sigma ef} \partial \omega_{\Sigma gh}} &= - \sum_{i=1}^n \frac{1}{4} (2 - \delta_{ef})(2 - \delta_{gh}) \\
&\times [\{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} \}_{eg} \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} \}_{hf} \\
&+ \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} \}_{eh} \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} \}_{gf}] \hat{\sigma}_{ML(-i)ef} \hat{\sigma}_{ML(-i)gh} \\
&+ \sum_{i=1}^n \frac{1}{4} (2 - \delta_{ef})(2 - \delta_{gh}) \sum_{(e,f)(g,h)}^2 \sum_{(e,f)(g,h)}^2 \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} \}_{eg} \\
&\times \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i) \}_h \\
&\times \{ (\boldsymbol{\Omega}_\Sigma \odot \hat{\boldsymbol{\Sigma}}_{ML(-i)})^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i) \}_f \hat{\sigma}_{ML(-i)ef} \hat{\sigma}_{ML(-i)gh} \\
(q^* \geq e \geq f \geq 1; q^* \geq g \geq h \geq 1).
\end{aligned}$$

S3.2 Multivariate multiple regression with normal errors when the covariance matrix of errors is given

$$\boldsymbol{\omega} = \text{vec}(\boldsymbol{\Omega}_B),$$

$$\begin{aligned}
L_{(-)} &\equiv \prod_{i=1}^n \frac{1}{(2\pi)^{q^*/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \{ -(\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' \\
&\quad \times \boldsymbol{\Sigma}^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i) / 2 \}, \\
D_{(-)} &= -\ln L_{(-)} = \frac{nq^*}{2} \ln(2\pi) + \frac{n}{2} \ln |\boldsymbol{\Sigma}| \\
&\quad + \frac{1}{2} \sum_{i=1}^n (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' \boldsymbol{\Sigma}^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i), \tag{S3.4}
\end{aligned}$$

$$\frac{\partial D_{(-)}}{\partial \boldsymbol{\Omega}_B} = - \sum_{i=1}^n \{ \boldsymbol{\Sigma}^{-1} (\mathbf{y}_i - \boldsymbol{\Omega}_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' \odot \hat{\mathbf{B}}_{ML(-i)},$$

giving

$$\sum_{i=1}^n \{ \Sigma^{-1} (\Omega_B \odot \hat{\mathbf{B}}_{ML(-i)}) \mathbf{x}_i \mathbf{x}_i' \} \odot \hat{\mathbf{B}}_{ML(-i)} = \sum_{i=1}^n (\Sigma^{-1} \mathbf{y}_i \mathbf{x}_i') \odot \hat{\mathbf{B}}_{ML(-i)}, \quad (S3.5)$$

which yields a simultaneous linear equation as $\text{Avec}(\Omega_B) = \mathbf{a}$, where

$$\begin{aligned} (\mathbf{A})_{(ab,cd)} &= (\Sigma^{-1})_{ac} \sum_{i=1}^n (\hat{\mathbf{B}}_{ML(-i)})_{cd} (\mathbf{x}_i \mathbf{x}_i')_{db} (\hat{\mathbf{B}}_{ML(-i)})_{ab}, \\ (\mathbf{a})_{(ab)} &= \sum_{i=1}^n (\Sigma^{-1} \mathbf{y}_i \mathbf{x}_i')_{ab} (\hat{\mathbf{B}}_{ML(-i)})_{ab} \\ &\quad (a, c = 1, \dots, q^*; b, d = 1, \dots, p^*) \end{aligned} \quad (S3.6)$$

where a double subscript notation for rows and columns is used. From (S3.6), $\text{vec}(\hat{\Omega}_B) = \mathbf{A}^{-1} \mathbf{a}$ is obtained.

S3.3 Multivariate multiple regression for non-normal errors or normal errors with $\Sigma = k \mathbf{I}_{(q^*)}$, where k is a positive constant and $\mathbf{I}_{(q^*)}$ is the $q^* \times q^*$ identity matrix

$$\boldsymbol{\omega} = \text{vec}(\Omega_B),$$

$$D_{(-)} = \sum_{i=1}^n (\mathbf{y}_i - \Omega_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i)' (\mathbf{y}_i - \Omega_B \odot \hat{\mathbf{B}}_{ML(-i)} \mathbf{x}_i), \quad (S3.7)$$

where

$$\hat{\mathbf{B}}_{(-i)} = \hat{\mathbf{B}}_{ML(-i)} = (\mathbf{X}_{(-i)}' \mathbf{X}_{(-i)})^{-1} \mathbf{X}_{(-i)}' \mathbf{Y}_{(-i)} (i = 1, \dots, n) \quad (S3.8)$$

is used as before,

$$\frac{\partial D_{(-)}}{\partial \Omega_B} = -2 \sum_{i=1}^n \{ (\mathbf{y}_i - \Omega_B \odot \hat{\mathbf{B}}_{(-i)} \mathbf{x}_i) \mathbf{x}_i' \} \odot \hat{\mathbf{B}}_{(-i)}. \quad (S3.9)$$

Then, we obtain

$$\sum_{i=1}^n \{ (\Omega_B \odot \hat{\mathbf{B}}_{(-i)}) \mathbf{x}_i \mathbf{x}_i' \} \odot \hat{\mathbf{B}}_{(-i)} = \sum_{i=1}^n (\mathbf{y}_i \mathbf{x}_i') \odot \hat{\mathbf{B}}_{(-i)}, \quad (S3.10)$$

which yields a simultaneous linear equation $\text{Avec}(\Omega_B) = \mathbf{a}$, where

$$\begin{aligned} (\mathbf{A})_{(ab,ac)} &= \sum_{i=1}^n (\hat{\mathbf{B}}_{(-i)})_{ac} (\mathbf{x}_i \mathbf{x}_i')_{cb} (\hat{\mathbf{B}}_{(-i)})_{ab}, \\ (\mathbf{a})_{(ab)} &= \sum_{i=1}^n (\mathbf{y}_i \mathbf{x}_i')_{ab} (\hat{\mathbf{B}}_{(-i)})_{ab} \quad (a = 1, \dots, q^*; b, c = 1, \dots, p^*), \end{aligned} \quad (S3.11)$$

which gives $\text{vec}(\hat{\Omega}_B) = \mathbf{A}^{-1}\mathbf{a}$.

S4. Poisson regression under canonical parametrization

$$\begin{aligned}
L_{(-)} &= \prod_{i=1}^n \exp\{-\exp(\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})\} \exp(\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)} y_i) / y_i!, \\
(y_i &= 0, 1, 2\dots), \\
D_{(-)} &= \sum_{i=1}^n \{\exp(\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)}) - \mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)} y_i + \ln(y_i!)\}, \\
\frac{\partial D_{(-)}}{\partial \boldsymbol{\omega}} &= \sum_{i=1}^n \{\exp(\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)}) - y_i\} \mathbf{x}_i \odot \hat{\beta}_{ML(-i)}, \\
\frac{\partial^2 D_{(-)}}{\partial \boldsymbol{\omega} \partial \boldsymbol{\omega}'} &= \sum_{i=1}^n \exp(\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)}) (\mathbf{x}_i \odot \hat{\beta}_{ML(-i)}) (\mathbf{x}_i \odot \hat{\beta}_{ML(-i)})'.
\end{aligned} \tag{S4.1}$$

S5. Logistic regression

$$\begin{aligned}
L_{(-)} &= \prod_{i=1}^n \left\{ \frac{1}{1 + \exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})} \right\}^{y_i} \left\{ \frac{\exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})}{1 + \exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})} \right\}^{1-y_i} \\
&= \prod_{i=1}^n \frac{\exp\{-(1-y_i)\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)}\}}{1 + \exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})}, \quad (y_i = 0, 1), \\
D_{(-)} &= \sum_{i=1}^n \left[(1-y_i) \mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)} + \ln\{1 + \exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})\} \right], \\
\frac{\partial D_{(-)}}{\partial \boldsymbol{\omega}} &= \sum_{i=1}^n \left\{ 1 - y_i - \frac{\exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})}{1 + \exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})} \right\} \mathbf{x}_i \odot \hat{\beta}_{ML(-i)} \\
&= \sum_{i=1}^n \left\{ \frac{1}{1 + \exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\beta}_{ML(-i)})} - y_i \right\} \mathbf{x}_i \odot \hat{\beta}_{ML(-i)}, \\
\frac{\partial^2 D_{(-)}}{\partial \boldsymbol{\omega} \partial \boldsymbol{\omega}'} &= \sum_{i=1}^n P_{(-i)}^{(\omega)} Q_{(-i)}^{(\omega)} (\mathbf{x}_i \odot \hat{\beta}_{ML(-i)}) (\mathbf{x}_i \odot \hat{\beta}_{ML(-i)})',
\end{aligned}$$

$$P_{(-i)}^{(\omega)} = \frac{1}{1 + \exp(-\mathbf{x}_i' \boldsymbol{\omega} \odot \hat{\boldsymbol{\beta}}_{ML(-i)})}, Q_{(-i)}^{(\omega)} = 1 - P_{(-i)}^{(\omega)} \quad (i = 1, \dots, n). \quad (S5.1)$$

Reference

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